

# Causal Inference with Continuous Multiple Time Point Interventions

Michael Schomaker

*Ludwig-Maximilians Universität München, Germany*

Joint Work with Iván Díaz, Paolo Denti, Helen McIlleron

**Talk at the Biometrical Colloqium, Lübeck**

28 February 2024

# Motivation – Continuous Interventions

## Data:

- ▶ CHAPAS-3 trial (📖 Mulenga et al., *Lancet Infectious Diseases*, 2016): children,  $\leq 13$  years, from Zambia/Uganda, randomized NRTI drug (abacavir, stavudine oder zidovudine) of HIV therapy

## Pharmacological Substudy:

- ▶ 📖 Bienczak et al. (*AIDS*, 2017) evaluated concentration of NNRTI drug regimen component: Nevirapine and Efavirenz  
→ higher probability of “viral failures” with lower concentrations
- ▶ what is the ideal target concentration?; causal question:

How many percent of children would have had a suppressed viral load at time  $t$  if they had had a concentration of “ $x$ ” mg/L efavirenz at each time point?

In general: how would probability of failure vary for different hypothetical concentration trajectories? → **“causal dose-response curve” (CDRC)**

# Motivation – Continuous Interventions (II)

We essentially have longitudinal observational data:

- ▶ Time-varying confounders: weight, adherence (with treatment-confounder feedback!)  
→ regression invalid!
  - ▶ Note: *Positivity assumption may not be satisfied with continuous interventions!*
  - ▶ Possible options to answer motivating question:
    - Option 1: G-methods → simple application (i.e., intervene for many trajectories) **today**
    - Option 2: Change question: “modified treatment policies” (e.g., Diaz et al., *JASA*, 2021)
    - Option 3: Find a compromise between interpretability and identifiability **today**
- Also: for 1 time point, great DR approach developed (Kennedy, *JRSS B*, 2017)

# Notation

- ▶ Follow-up time:  $t = 0, 1, \dots, T$
- ▶ Outcome:  $Y_t$
- ▶ Intervention:  $A_t$
- ▶ Confounder, Covariate:  $L_t$
- ▶ History: e.g.  $\bar{A}_t = (A_0, \dots, A_t)$
- ▶ History up to  $A_t$ :  $H_t$
- ▶ **Counterfactual:** e.g.  $Y_t^{\bar{a}_t}$

# Estimand & Estimation with Sequential G-computation

Estimand: Causal Dose-Response Curve

$$m_t : \bar{a}_t \mapsto E(Y_t^{\bar{a}_t} | \mathbf{L}_0^*), \quad t = 0, 1, \dots, T$$

Under sequential conditional exchangeability, consistency and **positivity** we have:

$$\mathbb{E}(Y_t^{\bar{a}_t}) = \mathbb{E}(\dots \mathbb{E}(\mathbb{E}(Y_t | \bar{\mathbf{A}}_t = \bar{a}_t, \mathbf{H}_t) | \bar{\mathbf{A}}_{t-1} = \bar{a}_{t-1}, \mathbf{H}_{t-1}) \dots | A_0 = a_0, \mathbf{L}_0)).$$

→ substitution estimation (sequential g-computation)

Positivity:

$$\inf_{a_t \in \mathcal{A}_t} g(a_t | \mathbf{h}_t) > 0 \quad \text{whenever} \quad p_0(\bar{\mathbf{L}}_t = \bar{\mathbf{I}}_t, \bar{\mathbf{A}}_{t-1} = \bar{a}_{t-1}) > 0 \quad \forall t, \bar{a}_t, \bar{\mathbf{I}}_t.$$

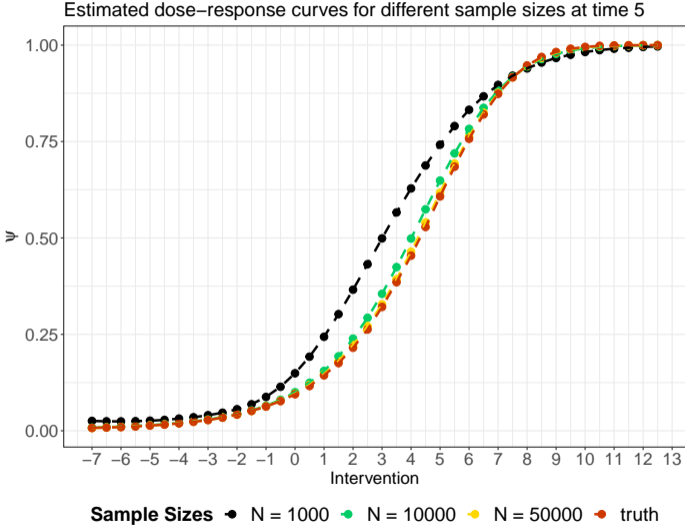
where  $\bar{\mathcal{A}}_t$  denotes the set of all relevant strategies  $\bar{a}_t = (a_0, \dots, a_t)$

→ What if we simply assume positivity and apply g-computation for many  $\bar{a}_t$ ?

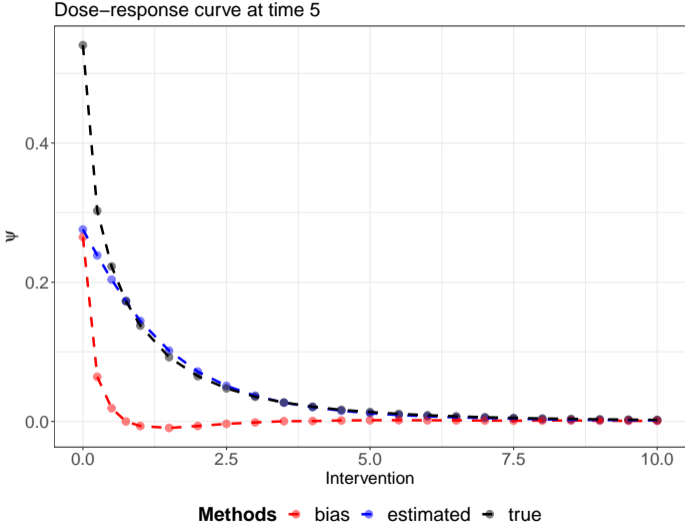
# Simulation (simple)



# Simulation (survival)



# Simulation (complex, as in data)





## Overall Consideration

The **tradeoff** to make is between

*estimating the CDRC as closely as possible, at the risk of bias due to positivity violations because of the continuous intervention*

and

*minimizing the risk of bias due to positivity violations, at the cost of redefining the estimand (e.g. by using modified treatment policies)*

Alternative: make a compromise!

## Proposal: Weighted Estimand (1 Time Point)

The general dose-response curve  $m : a \mapsto E(Y^a)$  can be identified<sup>1</sup> with the g-formula as

$$m(a) = \int E(Y \mid A = a, \mathbf{L} = \mathbf{l}) p_0(\mathbf{l}) d\nu(\mathbf{l}),$$

Proposal: instead, rather use

$$m_w(a) = \int E(Y \mid A = a, \mathbf{L} = \mathbf{l}) w(a, \mathbf{l}) p_0(\mathbf{l}) d\nu(\mathbf{l})$$

with

$$w(a, \mathbf{l}) = \begin{cases} 1 & \text{if } g(a \mid \mathbf{l}) > c \\ \frac{g(a \mid \mathbf{l})}{g(a)} & \text{otherwise.} \end{cases}$$

---

<sup>1</sup>under consistency, positivity and conditional exchangeability

## Weighted Estimand – Implications

- ▶ yields the desired dose-response curve under enough support (i.e.,  $g(a | \mathbf{I}) > c$ )
- ▶ otherwise the estimand is  $E(Y|A = a)$ 
  - not a causal quantity but does not require positivity assumption

## Weighted Estimand – Multiple Time Points

$$w_t(\mathbf{a}_{t+1}, \mathbf{h}_{t+1}, c) = \begin{cases} 1 & \text{if } g_t(\mathbf{a}_{t+1} \mid \mathbf{h}_{t+1}) > c, \\ \frac{g_t(\mathbf{a}_{t+1} \mid \mathbf{h}_{t+1})}{g_t(\mathbf{a}_{t+1} \mid \mathbf{a}_t, \mathbf{h}_t)} & \text{if } g_t(\mathbf{a}_{t+1} \mid \mathbf{h}_{t+1}) \leq c \text{ and } g_t(\mathbf{a}_{t+1} \mid \mathbf{a}_t, \mathbf{h}_t) > c, \\ \frac{g_t(\mathbf{a}_{t+1} \mid \mathbf{h}_{t+1})}{g_t(\mathbf{a}_{t+1} \mid \mathbf{a}_{t-1}, \mathbf{h}_{t-1})} & \text{if } g_t(\mathbf{a}_{t+1} \mid \mathbf{h}_{t+1}) \leq c \text{ and } g_t(\mathbf{a}_{t+1} \mid \mathbf{a}_t, \mathbf{h}_t) \leq c \\ & \text{and } g_t(\mathbf{a}_{t+1} \mid \mathbf{a}_{t-1}, \mathbf{h}_{t-1}) > c, \\ \vdots & \vdots \\ \frac{g_t(\mathbf{a}_{t+1} \mid \mathbf{h}_{t+1})}{g_t(\mathbf{a}_{t+1})} & \text{otherwise .} \end{cases}$$

## Weighted Estimand – Implications

- ▶ returns the CDRC if there is enough conditional support in terms of  $g_t(a_t | \mathbf{h}_t) > c$
- ▶ if there is not enough conditional support (and the weight denominator is  $> c$ )  
we can show that the estimand equates to  $E(Y_t | A_t = a_t, \dots, A_0 = a_0)$   
→ not a causal quantity but does not require positivity assumption
- ▶ if there is not enough conditional support and the weight denominator is too small the estimand entails a compromise

# Interpretation

We stick to the actual research question as long as possible, and calculate the CDRC in regions of enough support.

For some patients however, it may be unlikely (or even biologically impossible!) to actually observe some intervention trajectory of interest: those patients then “receive” *individual* concentration levels which generate outcomes that are typical for children with  $\bar{a}_t$  mg/l.<sup>2</sup>

For this, we make use of associations and require no positivity assumption.

The weighted curve acts like a magnifying glass and sensitivity tool if we don't want to rely on parametric extrapolation in regions of low support, where fixing the concentration to a specific level seems unrealistic.

---

<sup>2</sup>for example, children who are adherent to their drug regimen and got an appropriate drug dose prescribed, but are slow metabolizers will likely never be able to have very low concentration values.

## Weighted Estimand – Estimation (I)

For example, substitution estimator based on the following expression:

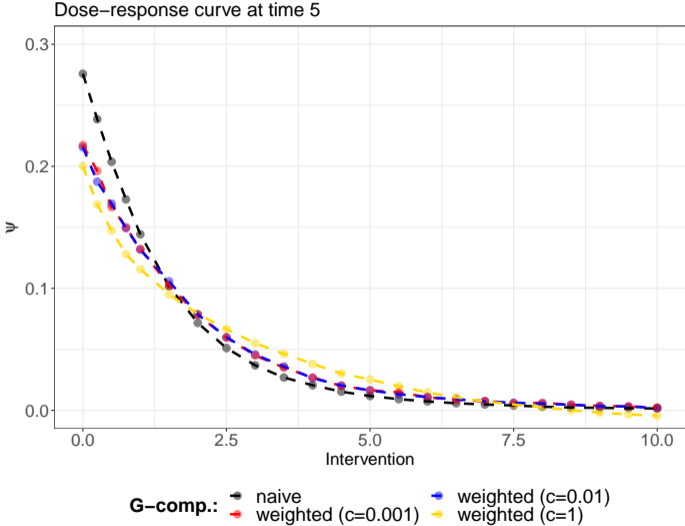
$$\mathbb{E}_{w_0}(\dots \mathbb{E}_{w_{t-1}}(\mathbb{E}_{w_t}(Y_t | \bar{A}_t = \bar{a}_t, \mathbf{H}_t) | \bar{A}_{t-1} = \bar{a}_{t-1}, \mathbf{H}_{t-1}) \dots | \mathbf{A}_0 = \mathbf{a}_0, \mathbf{L}_0),$$

where we define  $\mathbb{E}_{w_t}(Y_t | \bar{A}_t = \bar{a}_t, \mathbf{H}_t) = \mathbb{E}(w_t Y_t | \bar{A}_t = \bar{a}_t, \mathbf{H}_t)$ .

→ can also re-expressed into parametric g-formula-type expression, but then requires estimation of conditional *densities*, not only expectations

Note: even if  $Y_t$  is normal,  $w_t Y_t$  may not be normal; so we may need a data-adaptive approach

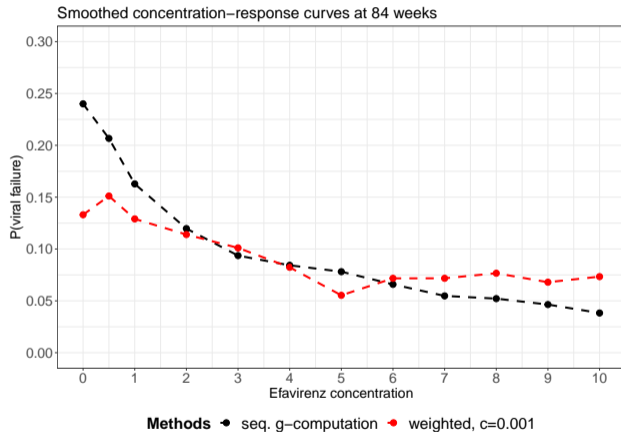
# Simulation (complex, as in data)





# Data Analysis

Based on a complete case analysis of  $n = 58$  kids

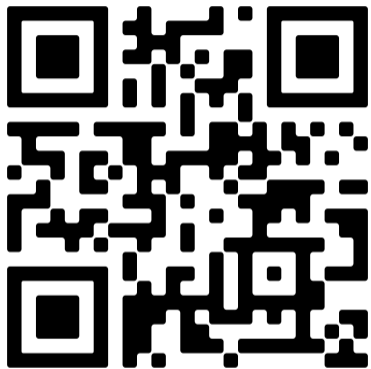


Weighted curve deviates from estimated CDRC in areas of low support

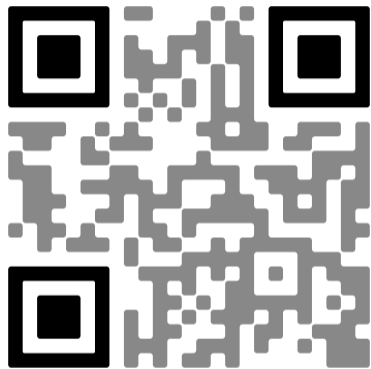
# Conclusions

- ▶ Standard g-computation can be used for continuous interventions
  - ▶ + targets the CDRC of interest
  - ▶ - relies on positivity assumption
- ▶ Simulations show that strategy may work, but can be problematic in regions of low support or with limited sample size
- ▶ Weighted curves offer a compromise, don't enforce unrealistic interventions, and do not require the positivity assumption
- ▶ The toolkit for causal effect estimation with longitudinal continuous interventions should ideally be broad

Read preprint on arXiv



...and test software



<https://arxiv.org/abs/2305.06645>

# Literature

- [1] Andrzej Bienczak, Paolo Denti, Adrian Cook, Lubbe Wiesner, Veronica Mulenga, Cissy Kityo, Addy Kekitiinwa, Diana M. Gibb, David Burger, Ann S. Walker, and Helen McIlleron.  
Determinants of virological outcome and adverse events in african children treated with paediatric nevirapine fixed-dose-combination tablets.  
*AIDS*, 31(7):905–915, 2017.
- [2] Iván Díaz, Nicholas Williams, Katherine L. Hoffman, and Edward J. Schenck.  
Non-parametric causal effects based on longitudinal modified treatment policies.  
*arXiv e-prints*, 2020.
- [3] Edward H. Kennedy, Zongming Ma, Matthew McHugh, and Dylan S. Small.  
Nonparametric methods for doubly robust estimation of continuous treatment effects.  
*Journal of the Royal Statistical Society. Series B, Statistical methodology*, 79:1229–1245, 2017.
- [4] Veronica Mulenga, Victor Musiime, Adeodata Kekitiinwa, Adrian D. Cook, George Abongomera, Julia Kenny, Chisala Chabala, Grace Mirembe, Alice Asiiimwe, Ellen Owen-Powell, David Burger, Helen McIlleron, Nigel Klein, Chifumbe Chintu, Margaret J. Thomason, Cissy Kityo, A. Sarah Walker, and Diana Gibb.  
Abacavir, zidovudine, or stavudine as paediatric tablets for african hiv-infected children (chapas-3): an open-label, parallel-group, randomised controlled trial.  
*The Lancet Infectious Diseases*, 16(2):169–79, 2016.

# APPENDIX

# Treatment-Confounder Feedback

